"Market efficiency and tax incentive policies during the COVID-19 pandemic: Case of Indonesia"

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MARKET EFFICIENCY AND TAX INCENTIVE POLICIES DURING THE COVID-19 PANDEMIC: CASE OF INDONESIA

Abstract

From an accounting perspective, taxes reduce profits and are often perceived as diminishing shareholders' rights. Consequently, government support through tax reductions plays a crucial role in enhancing the effectiveness of corporate strategies aimed at minimizing tax burdens. From the perspective of the Efficient Market Hypothesis (EMH), government tax incentive policies serve as vital signals to investors, shaping their expectations and influencing investment decisions. This study focuses on Indonesia's tax incentive policy introduced at the onset of the COVID-19 pandemic on April 1, 2020, and continuing until September 30, 2023. To assess market efficiency during this period, portfolios were constructed from the top 21 firms listed on the IDX Quality 30 and IDX High Dividend 20 indices, categorized by their systematic risk and cost of equity. The findings indicate that portfolios with higher systematic risk and cost of equity exhibit more optimal returns, greater volatility, and better risk-return tradeoffs. Conversely, portfolios with lower systematic risk and cost of equity tend to yield suboptimal returns due to their passive investment characteristics. Overall, the returns from all portfolios during the tax incentive period align with the weak form of the EMH, albeit showing negative autocorrelation instead of a purely random walk pattern. These findings imply that information regarding tax incentives influences prices primarily among firms with higher cost of equity or systematic risk. This study contributes to the understanding of the EMH by examining the impact of tax incentives during the pandemic while controlling for both the cost of equity and systematic risk.

Keywords

JEL Classification

return, systematic risk, equity cost, tax incentives, weak form, EMH, COVID-19

INTRODUCTION

Income taxes represent taxpayer obligations to the state. From an accounting perspective, taxes decrease the profits available to investors (Lintner, 1956; Modigliani & Miller, 1963). This suggests that lowering a firm's tax burden could potentially increase its profits, thereby enhancing its attractiveness to investors. Furthermore, the recent COVID-19 pandemic significantly impacted the global economy, with enduring repercussions that persist to the present. In response to the pandemic, the Indonesian Government implemented policies in 2020, including tax incentives, aimed at bolstering economic growth. Before the pandemic, business entities in Indonesia were subject to a 25% tax rate under Law No. 36 of 2008. From the perspective of the efficient market hypothesis (EMH), events such as the COVID-19 pandemic and the introduction of tax incentives provide significant information relevant to investor interests.

G11, G12, G14, G18, H25

Key events related to this tax policy began with the issuance of Government Regulation in Lieu of Law Number 1 on March 31, 2020, which was subsequently ratified in Law Number 2. Furthermore, on October 29, 2021, Law Number 7 of the Republic of Indonesia was is-

sued to harmonize tax regulations. On April 11, 2023, the Minister of Finance's Regulation Number 40 extended an income tax incentive of 3% for registered capital market firms. These developments were crucial for the performance of listed companies and were anticipated to contribute to national economic growth amid the COVID-19 pandemic. From April 1, 2020 to September 30, 2023, the stock market index exhibited an upward trajectory, indicating a positive investor response to the tax incentives. Consistent with the EMH, tax incentives likely captured investor attention, prompting movements in stock prices within the capital market. This assumes investors structured their portfolios optimistically, expecting reduced tax burdens to enhance firms' financial performance.

1. LITERATURE REVIEW

A market is deemed efficient when pertinent information is fully incorporated into current stock prices (Fama, 1970). Fama (1970) argued that such information reveals underlying risks, thereby influencing investors' expected returns. Malkiel (2003) contended that market risk is affected by unpredictable economic events. In testing for normal optimal returns, the weak form of the EMH, as proposed by Fama (1970), is employed, which posits that information relevant to current stock prices is independent of past prices.

Empirically, the weak-form EMH can elucidate how capital markets respond to information about economic events that influence firms' financial performance. For instance, Duarte-Duarte et al. (2014) discovered inefficiencies in the returns of several Colombian firms during the onset of the global economic crisis from 2008 to 2010. Conversely, Kok and Munir (2015) demonstrated that share prices in Malaysia's finance sector from 1997 to 2014 were efficient, driven by robust longterm prospects and resilience during financial crises. Almujamed (2018) observed that Qatar's stock market prices in various industrial sectors, such as consumer goods, services, industry, and insurance, reflected efficiency alongside economic growth from 2004 to 2017, supported by strong earnings per share. Heymans and Santana (2018) found that market indices comprising well-established firms in South Africa between July 3, 1997 and March 3, 2015, tended to be more efficient owing to their growth potential and solid fundamental strength. However, similar to findings by Nageri and Abdulkadir (2019) in Nigeria and Elangovan et al. (2022) in India, where economic growth was substantial, market inefficiencies persisted due to changes in investor behavior influenced by information asymmetry.

Since 2020, the efficiency of capital markets has been significantly influenced by the COVID-19 pandemic. For instance, Dias et al. (2020) observed that markets in the US, China, and Portugal exhibited inefficiencies compared to those in Spain, Ireland, Greece, Belgium, France, and Germany. Ozkan (2021) reported widespread market inefficiencies globally, particularly in the US, Spain, the UK, Italy, France, and Germany, leading to increased speculation and mispricing amid the pandemic. Similarly, Wang and Wang (2021) noted a decline in market efficiency in the US due to the pandemic, reflected in lower market returns. Dias et al. (2022) found that economic instability and investor pessimism stemming from the pandemic contributed to market inefficiencies, evident in regions such as Africa, the UK, Japan, and the US. Zebende et al. (2022) highlighted a passive return movement pattern in the US and G-20 countries, indicative of uncertain global economic conditions during the pandemic. In contrast, Bolek et al. (2022) suggested that Scandinavian markets became more efficient with increased investor optimism following information about the pandemic. They also found that the Baltic markets remained efficient, unaffected significantly by the COVID-19 pandemic (Bolek et al., 2022). Previous findings by He et al. (2020) in China indicated mixed impacts of the pandemic on the industry, not uniformly negative. Ammy-Driss and Garcin (2023) observed that US market conditions were less efficient compared to Asian and Australian markets during the pandemic crisis. Aslam et al. (2022) attributed changes in market efficiency worldwide, including in Asia and Europe in 2020, to herd behavior driven by fear and anxiety.

Studies by Rizvi et al. (2021), Athira and Ramesh (2023), Daly (2023), and Ispriyarso and Wibawa (2023) underscore the role of fiscal policies, such as tax incentives, in promoting economic recovery.

A government's tax incentive policy, which alleviates tax burdens, serves as crucial information in capital markets and significantly impacts investment returns. Takirambudde (1995) suggested that tax regulations with incentives play a crucial role in stimulating investment growth through capital markets. Chen (2015) demonstrated that effective utilization of tax incentives by listed firms can bolster profits, thereby influencing investor interest in the capital market. Goh et al. (2016) also highlighted that engaging in tax avoidance can alleviate the equity burden on shareholders. Moreover, Lazăr and Istrate (2018) affirmed that tax hikes diminish firm profitability. Drake et al. (2019) indicated that US investors respond positively to tax avoidance strategies as they enhance firm value, though Ahmad et al. (2023) found limited impact in Pakistan. Koch et al. (2023) discovered that tax incentives in the form of compensation for fiscal losses can enhance share price performance in capital markets. These findings align with Flagmeier et al. (2023), emphasizing that investors consider firm tax conditions as crucial disclosures influencing cost-benefit evaluations. However, Cao et al. (2023) cautioned that profit increases stemming from tax incentives may be short-lived.

Firms often employ various strategies to manage their tax obligations. For instance, in terms of capital structure, taxes are a primary motivator for firms to utilize debt as a tax shield (Jensen & Meckling, 1976; Myers, 2001; Fama & French, 2002; Frank & Goyal, 2003). However, the use of debt often correlates with the risk of bankruptcy (Nazir et al., 2021; Abdullah et al., 2022; Arhinful & Radmehr, 2023) and reduces funds available for investment (Thi et al., 2023). Delgado et al. (2018) observed that large German firms from 1992 to 2009 often utilized significant debt to create tax shields due to high tax burdens. Mocanu et al. (2021) found similar patterns among Romanian firms from 2013 to 2017. Dang and Tran (2021) illustrated how listed Vietnamese firms used debt to gain tax advantages from 2008 to 2020, noting that firms with substantial long-term assets engaged less in tax avoidance strategies. Recently, Lian (2022) revealed that increases in tax burdens prompt firms to allocate more resources to social and environmental initiatives, indirectly benefiting investors.

Based on the preceding empirical evidence, this study identifies three key points: First, economic conditions, whether favorable or adverse, are crucial in determining market efficiency. Second, market inefficiencies were more prevalent during the COVID-19 pandemic. Third, reducing the tax burden reflects the effectiveness of tax planning and attracts investor attention. The COVID-19 pandemic significantly impacted Indonesia's economic conditions, with tax incentives as a stimulus for firms to enhance financial performance and influence investor portfolio decisions. These factors played a key role in shaping market efficiency in its weak form. Thus, this study explores the relationship between systematic risk, cost of equity, and portfolio performance under tax incentives. It investigates their role in market efficiency during the COVID-19 pandemic in Indonesia, providing insights into how tax policies influenced this efficiency.

2. METHOD

This study focuses on the Indonesian market because of its unique economic context. Indonesia, as an emerging market in Southeast Asia, has demonstrated strong economic resilience since the COVID-19 pandemic (Siregar et al., 2021; Anas et al., 2022; Indrawati et al., 2024). Investors responded swiftly to the Indonesian government's economic strengthening efforts during the pandemic, particularly through tax incentives, which were perceived positively. This positivity was evident in the fluctuating but overall positive correction of the stock market index.

The study's sample consists of listed firms in Indonesia, specifically selected from two indices: IDX Quality 30 (IDXQ30) and IDX High Dividend 20 (IDXHIDIV20). The IDXQ30 comprises the top 30 firms based on strong financial performance indicators such as profitability, solvency, stable profit growth, and liquidity. The IDXHIDIV20 includes the top 20 firms with consistently high cash dividend distributions and yields. Firms included in the study had to maintain continuous inclusion in these indices throughout the observation period, from April 1, 2020 to September 30, 2023 (848 market days). The final sample consisted of 21 firms, resulting in 17,808 total observations over the observation period. April 1, 2020, was chosen as the starting point as it is the date of the initial announcement regarding tax incentives during the COVID-19 pandemic.

To test the hypothesis of market efficiency, the weak-form EMH is utilized. The methodology involves analyzing time-series data. First, R_{it} - RF_t is calculated, where R_{it} represents $(P_{it}-P_{it-1})/P_{it-1}$, and RF_t denotes the risk-free interest rate from the Central Bank of Indonesia at time *t*. Here, P_{it} represents the stock price of firm *i* at time *t*, while P_{it-1} represents the stock price at time *t*-1. Second, data normality is assessed using the Anderson-Darling (AD) test, as depicted in Equation (1).

$$AD = -N - \frac{1}{N} \sum_{i=1}^{N} (2i-1)$$

$$\times \left[(lnF(Y_i)) + ln(1 - F(Y_{N+1-i})) \right],$$
(1)

where *N* is the sample size, *F* is the cumulative distribution function, and Y_i is the *i*th ordered value of the sample. Third, determining the presence of a unit-root problem or stationary data is necessary using the Augmented Dickey-Fuller (ADF) test, as shown in Equation (2).

$$\Delta y_t = \alpha + Y_{yt-1} + V_t, \qquad (2)$$

where Δy_t is the change in the time series, yt-1 is the lagged value of the series, α is a constant term, and V_t is the error term. Fourth, the variance ratio (VR) test by Lo and MacKinlay (1988, 1989), as outlined in Equation (3), was applied to assess market efficiency within the framework of the EMH. Lo and MacKinlay (1988, 1989) extended Fama's (1970) methodology and established the VR test as a more robust approach for evaluating the EMH. This method's effectiveness has been supported by subsequent studies including Smith et al. (2002), Smith and Ryoo (2003), Borges (2011), Qu and Xiong (2019), Metghalchi et al. (2021), Dias et al. (2022), and Rönkkö et al. (2024).

$$VR(k) = \frac{\sigma^2(k)}{\sigma^2(1)},$$
(3)

where $\sigma^2(k)$ represents the variance of the k-period return defined in Equation (4):

$$\sigma^{2}(k) = \frac{1}{m} \sum_{t=1}^{T} \left(ln \frac{P_{t}}{P_{t-k}} - k \hat{\mu} \right)^{2}, \qquad (4)$$

where P_t and P_{t-k} are the asset prices at times *t* and *t*-*k*, $k\hat{\mu}$ is the expected return over the interval *k*, and *m* is defined as k(T-k+1)(1-k/T). *T* is the total number of observations, and *k* is the length of the interval. $\sigma^2(1)$ represents the variance of a period return (χ), defined as shown in Equation (5).

$$\sigma^{2}(1) = \frac{1}{T-1} \sum_{t=1}^{T} (\chi_{t} - \hat{\mu})^{2}, \qquad (5)$$

where χ_t is the observed return at time *t*, and $\hat{\mu}$ is the mean of the returns. The z-statistic for the variance ratio test is defined in Equation (6).

$$z(k) = \frac{VR(k) - 1}{\sqrt{\emptyset(k)}}.$$
(6)

Assuming homoscedasticity, $\phi(k)$ can be defined as in Equation (7).

$$\varnothing(k) = \frac{2(2k-1)(k-1)}{3kT},\tag{7}$$

where T represents the total number of observations, and k represents lags of 2, 4, 8, and 16. In the VR test, the null hypothesis is accepted if the probability exceeds the 10% significance level. To obtain specific results, R_{it} -RF_t is sorted based on systematic risk (β) and firm cost of equity (*ER*). This approach reflects the critical factors investors must consider when forming portfolios: systematic risk (Hundal et al., 2019; Hoesli & Johner, 2022; Liu, 2022) and the cost of equity (Athanasakou et al., 2020; Artikis & Kampouris, 2022; Abdollahi et al., 2023; Thien & Hung, 2023). β is classified as high risk (high β) if its value is 1 or higher; otherwise, it is considered low risk (low β). Similarly, *ER*, is classified as low (low *ER*) if its value is below the median, and high if it is above (high *ER*). β is derived using the capital asset pricing model (CAPM), as shown in Equation (8).

$$R_{it} - RF_t = \alpha_t + \beta_{RM_t - RF_t} + \mathcal{E}_t, \qquad (8)$$

where RM_t represents the market return at time *t*, calculated similarly to R_{it} . ER_i is estimated based on the CAPM, as detailed in Equation (9).

$$ER_{i} = RF_{t} + \beta \left(RM_{t} - RF_{t} \right). \tag{9}$$

3. RESULTS

Table 1 displays the descriptive statistics of stock returns adjusted for the risk-free rate (RT-RF). The findings indicate that portfolios based on high ER

and β exhibit higher average returns compared to portfolios with low ER and β . Specifically, the average ER for each portfolio is as follows: (1) RT-RF (high β) = 0.00066, (2) RT-RF (low β) = 0.00042, (3) RT-RF (high ER) = 0.00067, and (4) RT-RF (low ER) = 0.00043. These results suggest that portfolios with higher ER or β generally yield superior returns owing to their higher cost of equity. Furthermore, portfolios characterized by high ER and β show greater standard deviations, indicating greater volatility in market returns for these combinations. Additionally, the coefficient of variation (CV) reveals that portfolios with high ER offer a more favorable balance between risk and return compared to those with high β . The descriptive statistics also indicate that all portfolio combinations exhibit left-skewed distributions with positive (leptokurtic) kurtosis, suggesting that the sampled portfolios produced favorable returns.

Table 1. Descriptive statistics

RT-RF	Mean	SD	CV	Skew	Kurt
High β	0.000448	0.01349	3008.48	-0.16	2.86
Low β	0.000024	0.01033	43177.05	-0.53	5.84
High ER	0.000466	0.01397	2993.57	-0.19	2.78
Low ER	0.000047	0.01002	21528.34	-0.41	4.99

Note: RT-RF = risk-free rate; SD = standard deviation; CV = coefficient of variation; Kurt = kurtosis; β = risk; ER = cost of equity.

Figure 1 displays the results of testing the normality of RT-RF using the AD test. The null hypothesis of this test posits that the time-series data follow a normal distribution at a specified significance level. The AD test results indicate that the probabilities (*p*-values) for RT-RF are below 1%, 5%, and 10%. Consequently, based on these findings, the stock returns across the entire portfolio cannot be considered normally distributed during the tax incentive period.

Table 2 presents the results of the ADF test conducted on RT-RF to detect unit-root issues. Accepting the null hypothesis of the ADF test suggests that the time-series data exhibit a unit-root problem, indicating non-stationarity. The results of the ADF test reveal that all portfolios have tstatistics with probabilities below the 1%, 5%, and 10% significance levels. These findings imply that there is no unit-root problem in the RT-RF timeseries data for any of the portfolios; hence, the data are stationary.

Table 3 presents the results of the VR test conducted on RT-RF, categorized by high and low β and ER. The null hypothesis of the VR test posits that the time-series data follow a random walk or are efficient, depending on the chosen significance level. In this test, lags are employed to gauge the



Figure 1. Normality test

RT–RF	t-statistic					
	1% level	5% level	10% level	ADF test	Prob.	
High β	-3.437846	-2.864738	-2.568527	-31.15572	0.0000	
Low β	-3.437846	-2.864738	-2.568527	-32.57334	0.0000	
High ER	-3.437846	-2.864738	-2.568527	-31.09423	0.0000	
Low ER	-3.437846	-2.864738	-2.568527	-32.63563	0.0000	

Table 2. Unit root problem test

Note: Prob. = probability; β = risk; ER = cost of equity.

Table 3. Variance ratio test

RT–RF		In the America			
	Lag 2	Lag 4	Lag 8	Lag 16	- Joint test
		H	ligh β		
VR	0.9329	0.8797	0.9244	0.9448	
Stat.	-1.3806	-1.3503	-0.5541	-0.2801	1.3806
Prob.	0.1674	0.1769	0.5795	0.7794	0.5195
		L	.ow β		
VR	0.8888	0.8280	0.7989	0.7466	
Stat.	-1.9749	-1.7353	-1.4159	-1.2777	1.9749
Prob.	0.0483	0.0827	0.1568	0.2013	0.1796
		Н	igh ER		
VR	0.9349	0.8935	0.9446	0.9607	
Stat.	-1.3672	-1.2116	-0.4086	-0.2006	1.3672
Prob.	0.1716	0.2257	0.6828	0.8410	0.5290
		L	ow ER		
VR	0.8869	0.7970	0.7439	0.6920	
Stat.	-1.9663	-2.0214	-1.7845	-1.5342	2.0214
Prob.	0.0493	0.0432	0.0743	0.1250	0.1621

Note: β = risk; Prob. = probability; Stat. = statistical value; ER = cost of equity.

time taken for information to impact stock prices. Additionally, joint tests are utilized to affirm market efficiency across the entire period. Following Lo and MacKinlay (1988), a VR value equal to 1 under normal distribution signifies a pure random walk. Values greater or less than 1 indicate trending (positive autocorrelation) or mean-reverting (negative autocorrelation) behaviors, respectively.

In the joint test, a statistical value of 1.3806 with a probability of 0.5195 suggests that returns associated with high β portfolios tend to exhibit greater randomness over the observation period. Specifically, statistical values at lags 2, 4, 8, and 16 have probabilities above 1%, 5%, and 10%. These findings imply that market prices promptly reflect new information following tax incentive announcements, indicating a higher level of market efficiency. Figure 2 illustrates the trend in VR values for high β portfolios as presented in Table 3. The VR values for high β portfolios are 0.9329 (lag 2), 0.8797 (lag 4), 0.9244 (lag 8), and 0.9448 (lag 16), all below 1. These values suggest that returns tend to exhibit negative autocorrelation (mean-reverting behavior). Graphically, the VR values (depicted by blue lines) closely approach 1, indicating a rapid mean-reverting process.

In the joint test, the statistical value of 1.9749 with a probability of 0.1796 indicates that returns associated with low β portfolios tend to exhibit greater randomness over the observation period. Further evidence from individual tests on low β -based portfolios supports this conclusion. Specifically, statistical significance is observed at lags 2 and 4, at the 5% and 10% levels, respectively, suggesting market inefficiency in response to tax incentive announcements. However, statistical significance diminishes after lags 8 and 16, indicating a shift towards market efficiency in processing tax incentive information. These findings suggest that information regarding tax incentives is not promptly reflected in market prices. Additionally, the individual tests reveal that VR values at all lags (2, 4, 8,



VR stat. for cumulated RT-RF High ß with Robust +/- 2*S.E. Bands

Figure 2. Variance ratio trends of high β

and 16) are less than 1, indicating mean reversion. Figure 3 visually illustrates this with VR values at lags 2 and 4 being 0.8888 and 0.8280, respectively, close to 1, suggesting an accelerated mean-reverting process during market inefficiency or predictability. However, at lags 8 and 16, VR values decrease, indicating a slower mean-reverting process as market efficiency increases.

This study reveals that returns associated with high ER exhibit similar characteristics to those associated with high β . In the joint test, a statistical value of 1.3672 with a probability of 0.5290 indicates ef-

ficient market conditions for returns with high ER during the tax incentive period. Individual tests further substantiate these findings by showing that statistical values at lags 2, 4, 8, and 16 are insignificant, suggesting that stock market prices promptly incorporate information on tax incentives. Additionally, Figure 4 illustrates VR values across all lags, indicating symptoms of mean reversion, although they are near 1. This suggests that the market positively responds to tax incentive information, facilitating a faster mean-reverting process. Consequently, returns based on high ER demonstrate the highest mean and the lowest CV.



VR stat. for cumulated RT-RF Low ß with Robust +/- 2*S.E. Bands

Figure 3. Variance ratio trends of low β



VR stat. for cumulated RT-RF High ER with Robust +/- 2*S.E. Bands

Figure 4. Variance ratio trends of high ER



VR stat. for cumulated RT-RF Low ER with Robust +/- 2*S.E. Bands

Figure 5. Variance ratio trends of low ER

In the joint test, the test statistic of 2.0214 is not significant at the 1%, 5%, and 10% significance levels, indicating that we fail to reject the null hypothesis. This suggests that the market operates efficiently during the tax incentive announcement period, particularly for returns with low ER. In the individual tests, the statistical value only becomes insignificant at lag 16, indicating a delayed reflection of tax incentive information in stock market prices. Furthermore, VR values of 0.8869, 0.7970, 0.7439, and 0.6920 indicate mean-reverting returns as depicted in Figure 5. This decreasing pattern of VR values mirrors that of low β -based returns, suggesting a slower mean-reverting process.

4. DISCUSSION

This study examines the impact of tax incentives during the COVID-19 pandemic in Indonesia on market efficiency, focusing on systematic risk and cost of equity. The findings reveal that portfolios with high β and high ER generated higher average returns, albeit with greater volatility, consistent with the risk-return trade-off. Notably, high ER portfolios exhibited a more favorable risk-return balance than high β portfolios, highlighting the critical role of cost of equity in portfolio selection during economic uncertainty.

Furthermore, high β and high ER portfolios exhibit mean-reverting behavior, indicating that tax incentive announcements are efficiently incorporated into market prices, thereby enhancing market efficiency. These results align with Lintner (1956) and Modigliani & Miller (1963), who emphasize the role of tax incentives in shaping market adjustments through their effects on capital structure, cost of capital, and investment decisions. These findings reinforce the idea that policy interventions, such as tax incentives, influence investor expectations and return predictability, shaping broader market dynamics. In contrast, low β portfolios initially display market inefficiency, as returns at shorter lags (2 and 4) suggest delayed incorporation of tax incentive information. However, efficiency improves over longer lags (8 and 16), reflecting a gradual market adjustment. The accelerated mean reversion at shorter lags suggests initial overreaction, while the slower mean-reverting process at longer lags indicates increasing efficiency. Similarly, low ER portfolios demonstrate a delayed reflection of tax incentives in stock prices, with a pattern of declining VR values mirroring that of low β portfolios. Their lower mean returns and relatively high CV suggest a tendency toward passive investment behavior.

Overall, these findings align with Dias et al. (2020), He et al. (2020), Bolek et al. (2022), and Ammy-Driss and Garcin (2023) regarding the effects of the COVID-19 pandemic on market efficiency. Consistent with Fama (1970) and Malkiel (2003), the results indicate that tax incentive information is incorporated into stock prices, albeit with varying time lags, highlighting their role in enhancing market efficiency. Furthermore, in line with Takirambudde (1995), Chen (2015), Goh et al. (2016), Rizvi et al. (2021), Athira and Ramesh (2023), Daly (2023), Ispriyarso and Wibawa (2023), and Koch et al. (2023), tax incentives tend to elicit positive investor responses, particularly in more stable economic conditions.

CONCLUSION

This study examined the EMH within the context of government policies and investor behavior during the COVID-19 pandemic, specifically focusing on the period marked by the implementation of tax incentives in Indonesia. Rigorous methodologies were employed to test the hypotheses and ensure robust analytical outcomes. The findings indicate that portfolios characterized by higher equity costs and systematic risk tend to yield optimal returns. Specifically, the EMH was tested on four portfolios consisting of top stocks from the IDXQ30 and IDXHIDIV20 indices, with the results indicating weakform efficiency.

First, all portfolios exhibit mean-reverting returns throughout the tax incentive period, implying that returns tend to decline following information that contradicts market expectations. Second, despite not strictly adhering to a random walk, stocks with high systematic risk or cost of equity are viewed favorably by the market. Third, stocks with low systematic risk or cost of equity are met with a more passive market response to tax incentives, resulting in slower mean reversion. Overall, these findings suggest that the implementation of tax incentive policies sufficiently influences market efficiency for the top stocks listed on the IDXQ30 and IDXHIDIV20, considering both systematic risk and cost of equity. However, other attributes of these portfolios' return require further examination to determine if they conform to a random walk pattern. The study's scope is limited to Indonesia's top stocks during the early stages of the COVID-19 pandemic, emphasizing the need for further research to explore the EMH in countries with similar policies and to investigate additional return characteristics.

AUTHOR CONTRIBUTIONS

Conceptualization: Novi Swandari Budiarso. Data curation: Winston Pontoh. Formal analysis: Novi Swandari Budiarso. Funding acquisition: Winston Pontoh. Investigation: Winston Pontoh. Methodology: Novi Swandari Budiarso. Project administration: Winston Pontoh. Resources: Winston Pontoh. Software: Winston Pontoh. Supervision: Novi Swandari Budiarso. Validation: Winston Pontoh. Visualization: Winston Pontoh. Writing – original draft: Winston Pontoh. Writing – review & editing: Novi Swandari Budiarso.

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APPENDIX A

Table A1. Sample

Firms IDXQ30		IDXHIDIV20	Systematic risk (β)		Cost of equity (ER _i)	
	IDXQ30		High	Low	High	Low
ACES	V			V		V
ADRO		V	V		V	
ASII		V	V		V	
BBCA	V	V		V		V
BBNI		V	V		V	
BBRI		V	V		V	
BMRI	V	V	V		V	
CPIN	V		V		V	
HMSP	V	V		V		V
INDF		V		V		V
INTP	V		V		V	
ITMG		V		V		V
KLBF	V			V		V
MIKA	V			V		V
MNCN	V		V		V	
PTBA	V	V	V		V	
SCMA	V		V		V	
SIDO	V			V		V
TLKM	V	V	V			V
UNTR	V	V	V		V	
UNVR	V			V		V